

3. Sequential Estimation.

In section 2. we give out the MLE for mean of Gaussian R.V.

$$\mu_{ML}^{(N)} = \frac{1}{N} \sum_n x_n \quad \text{--- the estimation for } N \text{ samples}$$

$$= \frac{1}{N} X_N + \frac{N-1}{N} \mu_{ML}^{(N-1)} \quad \text{--- a weighted estimation involving } \mu_{ML}^{(N-1)}$$

$$= \mu_{ML}^{(N-1)} + \frac{1}{N} (X_N - \mu_{ML}^{(N-1)}) \quad \text{--- }$$

It seems that we can have a sequential approaching estimation for $\mu_{ML}^{(\infty)}$. However, this is not always the case, because the estimator may be very complicated. We need some other powerful method.

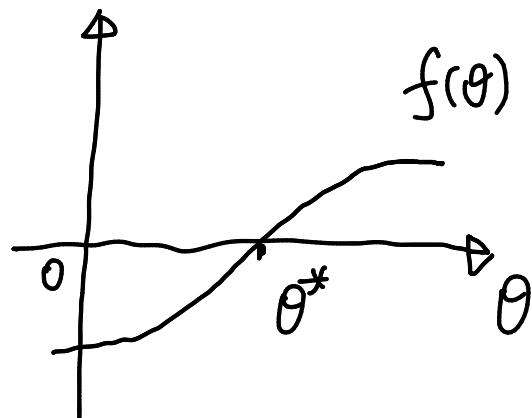
Robbins-Monro algorithm.

Given R.V.s $z \mid \theta$, $p(z|\theta)$

$$\text{def: } f(\theta) = \mathbb{E}(z|\theta) = \int z p(z|\theta) dz$$

$f(\theta)$ is called regression function.

Our goal is to find θ^* such that $f(\theta^*)=0$.



If we have a large dataset, we can turn to some models to estimate θ^* . But here we will get samples one by one. What we want is to make our estimated θ^* more and more accurate after each sample.

Assume the conditional variable of z is finite, so

$$\mathbb{E}[(z-f)^2 | \theta] < \infty$$

\uparrow
 $\mathbb{E}(z|\theta)$

also. w.l.o.g. we assume $f(\theta) > 0$ when $\theta > \theta^*$, $f(\theta) < 0$ when $\theta < \theta^*$

The Robbins-Monro procedure defines a sequential successive estimation of θ^* by

$$\theta^{(N)} = \theta^{(N-1)} + d_{N-1} z(\theta^{(N-1)})$$

an observation of z
when $\theta = \theta^{(N-1)}$

$\{d_N\}$ is a sequence of positive numbers with

$$\lim_{N \rightarrow \infty} d_N = 0 \quad \sum d_N = \infty \quad \sum d_N^2 < \infty \quad (\text{This is a reasonable step side})$$

Now let's see how this algorithm could solve MLE problem.

According to the setting.

$$\frac{\partial}{\partial \theta} \left\{ \frac{1}{N} \sum_{n=1}^N \ln p(x_n | \theta) \right\} \Big|_{\theta_{ML}} = 0$$

$$\lim_{N \rightarrow \infty} \frac{1}{N} \sum_{n=1}^N \frac{\partial}{\partial \theta} \ln p(x_n | \theta) \stackrel{\text{frequency} \mapsto \text{probability}}{\approx} \mathbb{E}_x \left[\frac{\partial}{\partial \theta} \ln p(x_n | \theta) \right]$$

then use

$$\theta^{(N)} = \theta^{(N-1)} + d_{N-1} \frac{\partial}{\partial \theta^{(N-1)}} \ln p(x_n | \theta^{(N-1)})$$

$$\text{for Gaussian R.V. } z = \frac{\partial}{\partial \mu_{ML}} \ln p(x | \mu_{ML}, \sigma^2) = \frac{1}{\sigma^2} (x - \mu_{ML})$$

If we choose $\alpha_N = \frac{\sigma^2}{N}$, we get

$$\begin{aligned}\theta^{(N)} &= \theta^{(N-1)} + \frac{\sigma^2}{N} \cdot \frac{1}{\sigma^2} \cdot (x - \mu_m) \\ &= \theta^{(N-1)} + \frac{1}{N} (x - \mu_m) \quad \text{-- same as previous result}\end{aligned}$$